Probabilistic Graphical Models

Learning with partial observations

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Fall 2019

Learning objectives

- different types of missing data
- learning with missing data and hidden vars:
 - directed models
 - undirected models
- develop an intuition for expectation maximization
 - variational interpretation

Two settings for partial observations

• missing data

• each instance in \mathcal{D} is missing some values

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latent variable models

- observations have common cause
- widely used in machine learning





image credit: Murphy's book

observation mechanism:

- generate the data point $X = [X_1, \ldots, X_D]$
- decide the values to observe $O_X = [1, 0, ..., 0, 1]$

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Learning with MCAR

missing completely at random (MCAR) P(X, O) = P(X)P(O)





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missing completely at random (MCAR)P(X,O) = P(X)P(O)headstails \swarrow \downarrow headstails $p(x) = \theta^x (1-\theta)^{1-x}$ throw to generate \checkmark \downarrow $p(o) = \psi^o (1-\theta)^{1-o}$ throw to decide show/hide



objective: learn a model for X, from the data $\mathcal{D} = \{x_o^{(1)}, \dots, x_o^{(M)}\}$ each x_o may include values for a different subset of vars.

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objective: learn a model for X, from the data $\mathcal{D} = \{x_o^{(1)}, \dots, x_o^{(M)}\}$ each x_o may include values for a different subset of vars.

since P(X,O) = P(X)P(O), we can ignore the obs. patterns

optimize: $\ell(\mathcal{D}, \theta) = \sum_{\boldsymbol{x_o} \in \mathcal{D}} \log \sum_{x_h} p(\boldsymbol{x_o}, x_h)$

A more general criteria

missing at random (MAR) $O_X \perp X_h | X_o$

if there is information about the obs. pattern O_X in X_h then it is also in X_o

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no "extra" information in the **obs. pattern** > ignore it

optimize: $\ell(\mathcal{D}, \theta) = \sum_{\boldsymbol{x}_o \in \mathcal{D}} \log \sum_{x_h} p(\boldsymbol{x}_o, x_h)$

marginal Likelihood function for partial observations

- fully observed data:
 - directed: likelihood decomposes
- ()• **undirected:** does not decompose, but it is concave

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marginal **Likelihood function** for partial observations

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- **undirected:** does not decompose, but it is concave
- partially observed:
- does not decompose
- ✓ not convex anymore

$$\ell(\mathcal{D}, heta) = \sum_{oldsymbol{x}_o \in \mathcal{D}} \log \sum_{x_h} p(oldsymbol{x}_o, x_h)$$

likelihood for a single assignment to the latent vars.



marginal Likelihood function: example

for a directed model

fully observed case decomposes:

 $\ell(D, heta) = \sum_{x,y,z\in\mathcal{D}}\log p(x,y,z)$

$$= \sum_x \log p(x) + \sum_{x,y} \log p(y|x) + \sum_{x,z} \log p(z|x)$$



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x is always missing (e.g., in a latent variable model)

 $\ell(D, heta) = \sum_{y, z \in \mathcal{D}} \log \sum_x p(x) p(y|x) p(z|x)$ cannot decompose it!

Parameter learning with missing data

Directed models:

option 1: obtain the gradient of marginal likelihood

option 2: expectation maximization (EM)

• variational interpretation

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• EM is not a good option here

all of these options need inference for each step of learning

example

log marginal likelihood:

$$\ell(\mathcal{D}) = \sum_{(a,d)\in\mathcal{D}} \log \sum_{m{b},m{c}} p(a) p(m{b}) p(m{c}|a,m{b}) p(d|m{c})$$



example

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take the derivative:

$$rac{\partial}{\partial p(d'|c')}\ell(\mathcal{D}) = rac{1}{p(d'|c')}\sum_{(a,d)\in\mathcal{D}} p(d',c'|a,d)$$
need inference for this



example

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need **inference** for this *what happens to this expression if every variable is observed?*

(directed models)

for a Bayesian Network with CPT

 $rac{\partial}{\partial p(x_i | pa_{x_i})} \ell(\mathcal{D}) = rac{1}{p(x_i | pa_{x_i})} \sum_{\mathbf{x}_o \in \mathcal{D}} p(x_i | pa_{x_i} | \mathbf{x}_o)$

some specific assignment

run inference for each observation

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a technical issue:

- gradient is always non-negative
 - lacksquare no constraint of the form $\sum_x p(x|pa_x) = 1$
 - reparametrize (e.g., using softmax)
 - ^O or use Lagrange multipliers

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for other parametrizations (beyond simple CPTs) use the chain rule:

$$rac{\partial}{\partial heta} \ell(\mathcal{D}; heta) = \sum_{(c', d') \in \mathcal{D}} rac{\partial \ell(\mathcal{D})}{\partial p(d' | c')} rac{\partial p(d' | c')}{\partial heta}$$



(directed models)

hidden

E-step: for each $a, d \in \mathcal{D}$ use the current parameters θ to get the marginals

more generally: expected sufficient statistics



A





in general we need inference to estimate this sufficient statistics



M-step:

use the marginals (similar to completely observed data) to learn θ expected sufficient statistics



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M-step:

use the marginals (similar to completely observed data) to learn θ expected sufficient statistics

E.g., update
$$\theta_{C|D}$$
 using $p_{\theta,\mathcal{D}}(C,D)$ and $p_{\theta,\mathcal{D}}(C)$ \longrightarrow $\theta_{D|C}^{new} = \frac{p_{\theta,\mathcal{D}}(C,D)}{p_{\theta,\mathcal{D}}(C)}$

(directed models)

for a Bayesian Network with CPT

E-step: for each $\mathbf{x}_o \in \mathcal{D}$ use the current parameters $\boldsymbol{\theta}$ to get the marginals

 $\{p_{\theta,\mathcal{D}}(X_i), p_{\theta,\mathcal{D}}(X_i, Pa_{X_i})\}$

(directed models)

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E-step: for each $\mathbf{x}_o \in \mathcal{D}$ use the current parameters $\boldsymbol{\theta}$ to get the marginals

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M-step:

use the marginals (similar to completely observed data) to learn θ^{new}

$$heta_{X_i|Pa_{X_i}}^{new} = rac{p_{ heta,\mathcal{D}}(X_i,Pa_{X_i})}{p_{ heta,\mathcal{D}}(Pa_{X_i})}$$
Expectation Maximization

(directed models)

for a Bayesian Network with CPT

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for undirected models: M-step is the expensive part

• perform E-step within each iteration of M-step: equivalent to gradient descent

Expectation Maximization: example



Expectation Maximization: example





Expectation Maximization: example

local optima in EM:



Expected log-likelihood

(directed models)

Original objective:

$$\ell(\mathcal{D}, heta) = \sum_{\mathbf{x}_o \in \mathcal{D}} rac{\log \sum_{\mathbf{x}_h} p_{ heta}(\mathbf{x}_o, \mathbf{x}_h)}{p_{ heta}(\mathbf{x}_o)}$$

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Recall: variational inference

 $\min_{q} \; \mathbb{D}_{KL}(q(\mathbf{x})|p(\mathbf{x})) = -\mathbb{H}(q) - \mathbb{E}_q[\log p(\mathbf{x})])$

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Recall: variational inference

- variational free energy

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for a latent variable model

 $\mathbb{D}_{KL}(q(\mathbf{x}_h)|p(\mathbf{x}_h|\mathbf{x}_o)) = -\mathbb{H}(q) - \mathbb{E}_q[\log p(\mathbf{x}_h,\mathbf{x}_o)] - \log p(\mathbf{x}_o)$

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 $\log p_{ heta}(\mathbf{x}_o) = \mathbb{D}_{KL}(q(\mathbf{x}_h)|p_{ heta}(\mathbf{x}_h|\mathbf{x}_o)) + \mathbb{H}(q) + \mathbb{E}_q[\log p_{ heta}(\mathbf{x}_h,\mathbf{x}_o)]$

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for a latent variable model

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expected log-likelihood wrt q
ignored by EM

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ignored by EM

Coordinate ascent:

- E-step: optimize q for a fixed θ (variational inference)
- M-step: optimize θ for a fixed q

EM as coordinate ascent

Coordinate ascent:

- E-step: optimize q for a fixed heta
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guaranteed improvement of $\log p_{ heta}(\mathbf{x}_o)$

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guaranteed improvement of $\log p_{ heta}(\mathbf{x}_o)$ converges to a local optimum

 $\log p_{ heta}(\mathbf{x}_o) = \mathbb{D}_{KL}(q(\mathbf{x}_h)|p_{ heta}(\mathbf{x}_h|\mathbf{x}_o)) + \mathbb{H}(q) + \mathbb{E}_q[\log p_{ heta}(\mathbf{x}_h,\mathbf{x}_o)]$

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evidence lower bound (ELBO) is a lower-bound on the likelihood

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 instead of $~q(\mathbf{x}_h)$

amortization: make q a *function* of observations

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 $q_\psi(\mathbf{x}_h \mid \mathbf{x}_o)$ instead of $q(\mathbf{x}_h)$ amortization: make q a *function* of observations $p_ heta$

$$p_{ heta}(\mathbf{x}_h,\mathbf{x}_o) = p_{ heta}(\mathbf{x}_h)p_{ heta}(\mathbf{x}_o|\mathbf{x}_h)$$

 $-\mathbb{D}_{KL}(q_{\psi}(\mathbf{x}_{h} \mid \mathbf{x}_{o})|p_{ heta}(\mathbf{x}_{h})) + \mathbb{E}_{q_{\psi}}[\log p_{ heta}(\mathbf{x}_{o}|\mathbf{x}_{h})]$

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Variational Auto-Encoder (VAE)

maximize ELBO by jointly optimizing $\psi, heta$ use neural networks to represent cond. distributions use back propagation for optimization

recall

linear exponential family

gradient in the fully observed setting

$$p(x; heta) = rac{1}{Z(heta)} \exp(\langle heta, \phi(x)
angle)$$

expectation wrt the model

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tion wrt the data expectation wrt the model

partial observation: $\mathbf{x} = (\mathbf{x}_o, \mathbf{x}_h)$ not observed

marginal likelihood: $p(\mathbf{x}_o; \theta) = \sum_{\mathbf{x}_h} \frac{1}{Z(\theta)} \exp(\langle \theta, \phi(\mathbf{x}) \rangle)$

recall

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gradient in the partially obs. case

$$abla_ heta\ell(heta,\mathcal{D}) = |\mathcal{D}| (rac{\mathbb{E}_{\mathcal{D}, heta}[\phi(x)]}{\downarrow} - \mathbb{E}_{p_ heta}[\phi(x)])
onumber \ \downarrow$$

wrt both data and model: we need to do inference to calculate expected sufficient statistics (similar to E-step in EM)

Example: Restricted Boltzmann Machine (RBM)

binary RBM:
$$p(h,v)=rac{1}{Z(heta)}\exp(\sum_{i,j} heta_{i,j}v_ih_j)$$

data: $\mathcal{D}=\{v^{(m)}\}_m$ for $v_i,h_j\in\{0,1\}$



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sufficient statistics: $\phi(v_i, h_j) = v_i h_j$
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we want to optimize: $\ell(\mathcal{D}; \theta) = \sum_{v \in \mathcal{D}} \log \sum_{h \in \mathcal{D}} \frac{1}{Z(\theta)} \exp(\sum_{i,j} \theta_{i,j} v_i h_j)$

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we want to optimize: $\ell(\mathcal{D}; \theta) = \sum_{v \in \mathcal{D}} \log \sum_{h \in \mathcal{D}} \log \sum_{i,j} \frac{1}{Z(\theta)} \exp(\sum_{i,j} \theta_{i,j} v_i h_j)$

 $\begin{array}{ll} \text{gradient:} & \frac{\partial}{\partial_{\theta_{i,j}}} \ell(\mathcal{D}; \theta) \propto \mathbb{E}_{\mathcal{D}, \theta}[v_i h_j] - \mathbb{E}_{p_{\theta}}[v_i h_j] \\ & = \left(\frac{1}{M} \sum_{v_i' \in \mathcal{D}} v_i' \mathbb{E}_{p_{\theta}}[h_j | v_i']\right) - \mathbb{E}_{p_{\theta}}[v_i h_j]) \end{array}$

Example: Restricted Boltzmann Machine (RBM)

binary RBM: $p(h,v)=rac{1}{Z(heta)}\exp(\sum_{i,j} heta_{i,j}v_ih_j)$ data: $\mathcal{D}=\{v^{(m)}\}_m$ for $v_i,h_j\in\{0,1\}$



sufficient statistics: $\phi(v_i, h_j) = v_i h_j$

we want to optimize: $\ell(\mathcal{D}; \theta) = \sum_{v \in \mathcal{D}} \log \sum_{h \in \mathcal{D}} \log \sum_{i,j} \exp(\sum_{i,j} \theta_{i,j} v_i h_j)$

 $\begin{array}{ll} \text{gradient:} & \frac{\partial}{\partial_{\theta_{i,j}}} \ell(\mathcal{D}; \theta) \propto \mathbb{E}_{\mathcal{D}, \theta}[v_i h_j] - \mathbb{E}_{p_{\theta}}[v_i h_j] \\ & = \left(\frac{1}{M} \sum_{v_i' \in \mathcal{D}} v_i' \mathbb{E}_{p_{\theta}}[h_j | v_i']\right) - \mathbb{E}_{p_{\theta}}[v_i h_j]) \end{array}$

sampling-based inference: sample h | v

use Gibbs sampling: sample both h,v using current parameters

summary

learning with partial observations:

- missing data
 - optimize the likelihood when missing at random
- latent variables
 - can produce expressive probabilistic models

problem is not convex how to learn the model?

- directly estimate the gradient (*directed and undirected*)
- use EM (directed models)
 - variational interpretation + relation to ELBO

$$\begin{array}{ccc} x & p(x;\pi) = \prod_k \pi_k^{\mathbb{I}(x=k)} \\ & & & & \\ & & & \\ & & & \\ & & & \\ & & & \\ \end{array} \\ \begin{array}{c} y \\ y \end{array} p(y|x;\{\mu_k,\Sigma_k\}) = \frac{1}{\sqrt{|2\pi\Sigma_x|}} \exp(-\frac{1}{2}(y-\mu_x)^T \Sigma_x^{-1}(y-\mu_x)) \end{array}$$

$$\begin{array}{c} \begin{array}{c} \begin{array}{c} x \\ \end{array} & p(x;\pi) = \prod_k \pi_k^{\mathbb{I}(x=k)} \\ & & & \\ \end{array} \\ \begin{array}{c} \text{model parameters } \end{array} \\ \theta = [\pi, \{\mu_k, \Sigma_k\}] \\ \downarrow \\ p(y|x; \{\mu_k, \Sigma_k\}) = \frac{1}{\sqrt{|2\pi\Sigma_x|}} \exp(-\frac{1}{2}(y-\mu_x)^T \Sigma_x^{-1}(y-\mu_x)) \end{array} \end{array}$$

E-step: calculate $\ p(x|y)$ for each $\ y \in \mathcal{D}$

$$p(x|y) \propto p(x;\pi) p(y|x;\mu,\Sigma) \ = \pi_k \mathcal{N}(y;\mu_k,\Sigma_k)$$

- now we have "probabilistically completed" instances
- update the parameters (easy in a Bayes-net)

$$\begin{array}{ccc} x & p(x;\pi) = \prod_k \pi_k^{\mathbb{I}(x=k)} \\ & & & & \\ & & & & \\ & & & \\ & & & & \\ & & & \\ & & & \\ & & & \\ & & & & & \\ & & & & \\ & & & & \\ & & & & & \\ & & & & & \\ & & & & &$$

M-step: estimate $\pi, \mu_k, \Sigma_k \forall k$

$$\begin{aligned} \pi_k^{new} &= \frac{1}{N} \sum_{y \in \mathcal{D}} \frac{p(x=k|y)}{\sum_{k'} p(x=k'|y)} \\ \mu_k &= \frac{\sum_{y \in \mathcal{D}} p(x=k|y)y}{\sum_{y \in \mathcal{D}} p(x=k|y)} \quad \text{mean of a weighted set of instances} \\ \\ \Sigma_k &= \frac{\sum_{y \in \mathcal{D}} p(x=k|y)(y-\mu_k)(y-\mu_k)^T}{\sum_{y \in \mathcal{D}} p(x=k|y)} \quad \text{covariance of a weighted set of instances} \end{aligned}$$

$$\begin{array}{c} \begin{array}{c} \begin{array}{c} x \\ y \end{array} & p(x;\pi) = \prod_k \pi_k^{\mathbb{I}(x=k)} \\ & & \\ & & \\ \end{array} \\ \begin{array}{c} y \end{array} & p(y|x;\{\mu_k,\Sigma_k\}) = \frac{1}{\sqrt{|2\pi\Sigma_x|}} \exp(-\frac{1}{2}(y-\mu_x)^T\Sigma_x^{-1}(y-\mu_x)) \end{array} \end{array}$$

M-step: estimate $\pi, \mu_k, \Sigma_k \forall k$

$$\begin{split} \pi_k^{new} &= \frac{1}{N} \sum_{y \in \mathcal{D}} \frac{p(x=k|y)}{\sum_{k'} p(x=k'|y)} \\ \mu_k &= \frac{\sum_{y \in \mathcal{D}} p(x=k|y)y}{\sum_{y \in \mathcal{D}} p(x=k|y)} \quad \text{mean of a weighted set of instances} \\ \Sigma_k &= \frac{\sum_{y \in \mathcal{D}} p(x=k|y)(y-\mu_k)(y-\mu_k)^T}{\sum_{y \in \mathcal{D}} p(x=k|y)} \quad \text{covariance of a weighted set of instances} \end{split}$$

