Applied Machine Learning

Regularization

Siamak Ravanbakhsh

COMP 551 (winter 2020)

1

Learning objectives

Basic idea of

- overfitting and underfitting
- Regularization (L1 & L2)
- MLE vs MAP estimation
- bias and variance trade off
- evaluation metrics & cross validation

Previously...

Linear regression and logistic regression is linear too simple? what if it's not a good fit? how to increase the models expressiveness?

- create new nonlinear features
- is there a downside?

replace original features in $\ f_w(x) = \sum_d w_d x_d$

replace original features in $\ f_w(x) = \sum_d w_d x_d$ with nonlinear bases $\ f_w(x) = \sum_d w_d \, \phi_d(x)$

replace original features in $\ f_w(x) = \sum_d w_d x_d$ with nonlinear bases $\ f_w(x) = \sum_d w_d \, \phi_d(x)$ linear least squares solution $\ w^* = (\Phi^{ op} \Phi)^{-1} \Phi^{ op} y$ replacing X with Φ

$$\Phi = egin{bmatrix} \phi_1(x^{(1)}), & \phi_2(x^{(1)}), & \cdots, & \phi_D(x^{(1)}) \ \phi_1(x^{(2)}), & \phi_2(x^{(2)}), & \cdots, & \phi_D(x^{(2)}) \ dots & dots & \ddots & dots \ \phi_1(x^{(N)}), & \phi_2(x^{(N)}), & \cdots, & \phi_D(x^{(N)}) \end{bmatrix}$$

replace original features in $\ f_w(x) = \sum_d w_d x_d$ with nonlinear bases $f_w(x) = \sum_d w_d \, \phi_d(x)$ linear least squares solution $w^* = (\Phi^ op \Phi)^{-1} \Phi^ op y$ replacing X with Φ a (nonlinear) feature $\Phi = \begin{bmatrix} \phi_1(x^{(1)}), & \phi_2(x^{(1)}), & \cdots, & \phi_D(x^{(1)}) \\ \phi_1(x^{(2)}), & \phi_2(x^{(2)}), & \cdots, & \phi_D(x^{(2)}) \\ \vdots & \vdots & \ddots & \vdots \\ \phi_1(x^{(N)}), & \phi_2(x^{(N)}), & \cdots, & \phi_D(x^{(N)}) \end{bmatrix}$ one instance

original input is scalar $x\in\mathbb{R}$ examples 0.75 0.75 0.5 0.5 0 0.5 -0.5 0.25 0.25 0 -10 0 0 -10 -1 -1 Sigmoid bases Gaussian bases polynomial bases $\phi_k(x)=e^{-rac{(x-\mu_k)^2}{s^2}}$ $\phi_k(x)=rac{1}{1+e^{-rac{x-\mu_k}{s}}}$ $\phi_k(x)=x^k$





$$\phi_k(x)=e^{-rac{(x-\mu_k)^2}{s^2}}$$







our fit to data using 10 Gaussian bases

$$\phi_k(x)=e^{-rac{(x-\mu_k)^2}{s^2}}$$



our fit to data using 10 Gaussian bases

prediction for a new instance

$$f(x') = \phi(x')^{ op} (\Phi^{ op} \Phi)^{-1} \Phi^{ op} y$$

| w found using LLS
features evaluated for the new point









Generalization



Overfitting



Overfitting



Overfitting



Model selection

how to pick the model with lowest expected loss / test error?

regularization bound the test error by bounding

- training error
- model complexity

Model selection

how to pick the model with lowest expected loss / test error?

regularization bound the test error by bounding

- training error
- model complexity

USE a validation set (and a separate test set for final assessment)



Model selection

how to pick the model with lowest expected loss / test error?

regularization bound the test error by bounding

- training error
- model complexity



USE a validation set (and a separate test set for final assessment)



when overfitting, we often see large weights



when overfitting, we often see large weights





when overfitting, we often see large weights





when overfitting, we often see large weights





when overfitting, we often see large weights



dashed lines are $\, w_d \phi_d(x) \quad orall d\,$



idea: penalize large parameter values

L2 regularized linear least squares regression:

$$J(w) = rac{1}{2} ||Xw - y||_2^2 + rac{\lambda}{2} ||w||_2^2$$

L2 regularized linear least squares regression:

$$J(w) = rac{1}{2}||Xw-y||_2^2 + rac{\lambda}{2}||w||_2^2$$

 $||w||_2^2$ sum of squared error $rac{1}{2}\sum_n(y^{(n)}-w^ op x)^2$

L2 regularized linear least squares regression:

$$egin{aligned} J(w) &= rac{1}{2}||Xw-y||_2^2 + rac{\lambda}{2}||w||_2^2 \ &ert$$
 sum of squared error (squared) L2 norm of w $rac{1}{2}\sum_n(y^{(n)}-w^ op x)^2$ $w^Tw &= \sum_d w^2 \end{aligned}$

L2 regularized linear least squares regression:

$$egin{aligned} J(w) &= rac{1}{2}||Xw-y||_2^2 + rac{\lambda}{2}||w||_2^2 \ &ert$$
 sum of squared error (squared) L2 norm of w $rac{1}{2}\sum_n(y^{(n)}-w^ op x)^2$ $w^Tw &= \sum_d w^2 \end{aligned}$

regularization parameter $\lambda > 0$ controls the strength of regularization

L2 regularized linear least squares regression:

$$egin{aligned} J(w) &= rac{1}{2}||Xw-y||_2^2 + rac{\lambda}{2}||w||_2^2 \ &ert$$
 sum of squared error (squared) L2 norm of w $rac{1}{2}\sum_n(y^{(n)}-w^ op x)^2$ $w^Tw &= \sum_d w^2 \end{aligned}$

regularization parameter $\lambda > 0$ controls the strength of regularization a good practice is to not penalize the intercept $\lambda(||w||_2^2 - w_0^2)$

we can set the derivative to zero $J(w) = \frac{1}{2}(Xw - y)^{\top}(Xw - y) + \frac{\lambda}{2}w^{\top}w$ $\nabla J(w) = X^{\top}(Xw - y) + \lambda w = 0$

we can set the derivative to zero $J(w) = \frac{1}{2}(Xw - y)^{\top}(Xw - y) + \frac{\lambda}{2}w^{\top}w$ $\nabla J(w) = X^{\top}(Xw - y) + \frac{\lambda}{2}w = 0$

> when using gradient descent, this term reduces the weights at each step (weight decay)

we can set the derivative to zero $J(w) = \frac{1}{2}(Xw - y)^{\top}(Xw - y) + \frac{\lambda}{2}w^{\top}w$ $\nabla J(w) = X^{\top}(Xw - y) + \frac{\lambda}{2}w = 0$

 $(X^ op X + \lambda \mathbf{I})w = X^ op y$

when using gradient descent, this term reduces the weights at each step **(weight decay)**
Ridge regression

we can set the derivative to zero $J(w) = \frac{1}{2}(Xw - y)^{\top}(Xw - y) + \frac{\lambda}{2}w^{\top}w$ $\nabla J(w) = X^{\top}(Xw - y) + \frac{\lambda}{2}w^{\top}w$

 $(X^ op X + \lambda \mathbf{I})w = X^ op y$

when using gradient descent, this term reduces the weights at each step **(weight decay)**

 $w = (X^ op X + \lambda \mathbf{I})^{-1} X^ op y$

Ridge regression

we can set the derivative to zero $J(w) = \frac{1}{2}(Xw - y)^{\top}(Xw - y) + \frac{\lambda}{2}w^{\top}w$ $\nabla J(w) = X^{\top}(Xw - y) + \frac{\lambda}{2}w^{\top}w$

 $(X^ op X + \lambda \mathbf{I})w = X^ op y$

when using gradient descent, this term reduces the weights at each step **(weight decay)**

$$w = (X^ op X + \lambda \mathbf{I})^{-1} X^ op y$$

the only part different due to regularization

Ridge regression

we can set the derivative to zero $J(w) = \frac{1}{2}(Xw - y)^{\top}(Xw - y) + \frac{\lambda}{2}w^{\top}w$ $\nabla J(w) = X^{\top}(Xw - y) + \frac{\lambda}{2}w = 0$

 $(X^ op X + \lambda \mathbf{I})w = X^ op y$

when using gradient descent, this term reduces the weights at each step **(weight decay)**

the only part different due to regularization

 λI makes it invertible!

we can have linearly dependent features (e.g., D > N)

 $w = (X^ op X + \lambda \mathbf{I})^{-1} X^ op y$

the solution will be unique!



polynomial bases

 $\phi_k(x)=x^k$

Without regularization:



polynomial bases

$$\phi_k(x)=x^k$$

Without regularization:





polynomial bases

$$\phi_k(x)=x^k$$

Without regularization:





polynomial bases

$$\phi_k(x)=x^k$$

Without regularization:





polynomial bases

 $\phi_k(x)=x^k$

with regularization:



polynomial bases

$$\phi_k(x)=x^k$$

with regularization:





polynomial bases

$$\phi_k(x)=x^k$$

with regularization:





polynomial bases

$$\phi_k(x)=x^k$$

with regularization:



what if we scale the input features, using different factors $\tilde{x}^{(n)} = \gamma_d x^{(n)} orall d, n$

what if we scale the input features, using different factors $\tilde{x}^{(n)} = \gamma_d x^{(n)} orall d, n$

if we have no regularization: $ilde w_d = rac{1}{\gamma_d} w_d orall d$ everything remains the same because: $||Xw-y||_2^2 = || ilde X ilde w - y||_2^2$

what if we scale the input features, using different factors $\tilde{x}^{(n)} = \gamma_d x^{(n)} orall d, n$

if we have no regularization: $ilde w_d = rac{1}{\gamma_d} w_d orall d$ everything remains the same because: $||Xw-y||_2^2 = || ilde X ilde w - y||_2^2$

with regularization: $||\tilde{w}||_2 \neq ||w||_2^2$ so the optimal **w** will be different!

what if we scale the input features, using different factors $\tilde{x}^{(n)} = \gamma_d x^{(n)} orall d, n$

if we have no regularization: $ilde w_d = rac{1}{\gamma_d} w_d orall d$ everything remains the same because: $||Xw-y||_2^2 = || ilde X ilde w - y||_2^2$

with regularization: $||\tilde{w}||_2 \neq ||w||_2^2$ so the optimal **w** will be different!

features of different mean and variance will be penalized differently

normalization

$$egin{aligned} \mu_d &= rac{1}{N} x_d^{(n)} \ \sigma_d^2 &= rac{1}{N-1} (x_d^{(n)} - \mu_d)^2 \end{aligned}$$

makes sure all features have the same mean and variance $\, x_d^{(n)} \,$

$$egin{array}{ccc} x_d^{(n)} - \mu_d \ \overline{\sigma_d} \end{array}$$

previously: linear regression & logistic regression maximize log-likelihood

previously: linear regression & logistic regression maximize log-likelihood

linear regression

$$w^* = rg \max p(y|w)$$

$$egin{aligned} &= rg\max_w \prod_{n=1}^N \mathcal{N}ig(y;\Phi w,\sigma^2ig) \ &\equiv rg\min\sum_n L_2(y^{(n)},w^ opig\phi(x^{(n)})) \end{aligned}$$

previously: linear regression & logistic regression maximize log-likelihood

previously: linear regression & logistic regression maximize log-likelihood

idea: maximize the posterior instead of likelihood

$$p(w|y) = rac{p(w)p(y|w)}{p(y)}$$

use the Bayes rule and find the parameters with max posterior prob.

$$p(w|y) = rac{p(w)p(y|w)}{p(y)}$$
 the same for all choices of w (ignore)

use the Bayes rule and find the parameters with max posterior prob.

$$p(w|y) = rac{p(w)p(y|w)}{p(y)}$$
 the same for all choices of w (ignore)

MAP estimate

 $w^* = rg\max_w p(w) p(y|w)$

use the Bayes rule and find the parameters with max posterior prob.

$$p(w|y) = rac{p(w)p(y|w)}{p(y)}$$
 the same for all choices of w (ignore)

MAP estimate

 $w^* = rg\max_w p(w) p(y|w)$

 $\equiv rg\max_w \log p(y|w) + \log p(w)$

use the Bayes rule and find the parameters with max posterior prob.

$$p(w|y) = rac{p(w)p(y|w)}{p(y)}$$
 the same for all choices of w (ignore)

MAP estimate

 $w^* = rg\max_w p(w) p(y|w)$

 $\equiv \arg \max_{w} \log p(y|w) + \log p(w)$ likelihood: original objective

use the Bayes rule and find the parameters with max posterior prob.

$$p(w|y) = rac{p(w)p(y|w)}{p(y)}$$
 the same for all choices of w (ignore)

MAP estimate

 $w^* = rg\max_w p(w) p(y|w)$

 $\equiv \arg \max_{w} \log p(y|w) + \frac{\log p(w)}{\operatorname{prior}}$

use the Bayes rule and find the parameters with max posterior prob.

$$p(w|y) = rac{p(w)p(y|w)}{p(y)}$$
 the same for all choices of w (ignore)

MAP estimate

 $w^* = rg\max_w p(w) p(y|w)$

 $\equiv \arg \max_{w} \log p(y|w) + \frac{\log p(w)}{\operatorname{prior}}$

even better would be to estimate the posterior distribution p(w|y)

• more on this later in the course!

Gaussian likelihood and Gaussian prior

 $w^* = rg\max_w p(w) p(y|w)$

Gaussian likelihood and Gaussian prior

 $w^* = rg\max_w p(w) p(y|w) ~~\equiv rg\max_w \log p(y|w) + \log p(w)$

Gaussian likelihood and Gaussian prior

 $w^* = rg \max_w p(w) p(y|w) ~~\equiv rg \max_w \log p(y|w) + \log p(w)$

 $\equiv rg\max_w \log \mathcal{N}(y|w^ op x, \sigma^2) + \sum_{d=1}^D \log \mathcal{N}(w_d, 0, au^2)$

Gaussian likelihood and Gaussian prior

 $w^* = rg\max_w p(w) p(y|w) ~~\equiv rg\max_w \log p(y|w) + \log p(w)$

 $\equiv \arg \max_{w} \log \mathcal{N}(y|w^{\top}x, \sigma^{2}) + \sum_{d=1}^{D} \log \mathcal{N}(w_{d}, 0, \tau^{2})$ assuming independent Gaussian (one per each weight)

Gaussian likelihood and Gaussian prior

 $w^* = rg\max_w p(w) p(y|w) ~~\equiv rg\max_w \log p(y|w) + \log p(w)$

 $\equiv \arg \max_{w} \log \mathcal{N}(y|w^{\top}x, \sigma^{2}) + \sum_{d=1}^{D} \log \mathcal{N}(w_{d}, 0, \tau^{2})$ assuming independent Gaussian (one per each weight)

 $\equiv rg\max_w rac{-1}{2\sigma^2}(y-w^ op x)^2 - \sum_{d=1}^D rac{1}{2 au^2}w_d^2$

Gaussian likelihood and Gaussian prior

 $w^* = rg\max_w p(w) p(y|w) ~~\equiv rg\max_w \log p(y|w) + \log p(w)$

 $\equiv \arg \max_{w} \log \mathcal{N}(y|w^{\top}x, \sigma^{2}) + \sum_{d=1}^{D} \log \mathcal{N}(w_{d}, 0, \tau^{2})$ assuming independent Gaussian (one per each weight)

 $x \equiv rg \max_w rac{-1}{2\sigma^2} (y - w^ op x)^2 - \sum_{d=1}^D rac{1}{2 au^2} w_d^2 \;\; \equiv rg \min_w rac{1}{2} (y - w^ op x)^2 + \sum_{d=1}^D rac{\sigma^2}{2 au^2} w_d^2$

Gaussian likelihood and Gaussian prior

 $w^* = rg\max_w p(w) p(y|w) ~~\equiv rg\max_w \log p(y|w) + \log p(w)$

 $\equiv \arg \max_{w} \log \mathcal{N}(y|w^{\top}x, \sigma^{2}) + \sum_{d=1}^{D} \log \mathcal{N}(w_{d}, 0, \tau^{2})$ assuming independent Gaussian (one per each weight)

 $x \equiv rg \max_w rac{-1}{2\sigma^2} (y - w^ op x)^2 - \sum_{d=1}^D rac{1}{2 au^2} w_d^2 \;\; \equiv rg \min_w rac{1}{2} (y - w^ op x)^2 + \sum_{d=1}^D rac{\sigma^2}{2 au^2} w_d^2$

multiple data-points

$$\equiv rgmin_w rac{1}{2} \sum_n (y^{(n)} - w^ op x^{(n)})^2 + \sum_{d=1}^D rac{\lambda}{2} w_d^2$$

Gaussian likelihood and Gaussian prior

 $w^* = rg\max_w p(w) p(y|w) ~~\equiv rg\max_w \log p(y|w) + \log p(w)$

 $\equiv \arg \max_{w} \log \mathcal{N}(y|w^{\top}x, \sigma^{2}) + \sum_{d=1}^{D} \log \mathcal{N}(w_{d}, 0, \tau^{2})$ assuming independent Gaussian (one per each weight)

 $x \equiv rg \max_w rac{-1}{2\sigma^2} (y - w^ op x)^2 - \sum_{d=1}^D rac{1}{2 au^2} w_d^2 \ \ \equiv rg \min_w rac{1}{2} (y - w^ op x)^2 + \sum_{d=1}^D rac{\sigma^2}{2 au^2} w_d^2$

multiple data-points $= \arg \min_{w} \frac{1}{2} \sum_{n} (y^{(n)} - w^{\top} x^{(n)})^{2} + \sum_{d=1}^{D} \frac{\lambda}{2} w_{d}^{2}$ L2 regularization

Gaussian likelihood and Gaussian prior

 $w^* = rg\max_w p(w) p(y|w) ~~\equiv rg\max_w \log p(y|w) + \log p(w)$

 $\equiv \arg \max_{w} \log \mathcal{N}(y|w^{\top}x, \sigma^{2}) + \sum_{d=1}^{D} \log \mathcal{N}(w_{d}, 0, \tau^{2})$ assuming independent Gaussian (one per each weight)

L2- regularization is assuming a Gaussian prior on weights the same is true for logistic regression (or any other cost function)

Laplace prior

another notable choice of prior is the Laplace distribution

Laplace prior

another notable choice of prior is the Laplace distribution



image:https://stats.stackexchange.com/questions/177210/why-is-laplace-prior-producing-sparse-solutions
Laplace prior

another notable choice of prior is the Laplace distribution

minimizing negative log-likelihood $ightarrow -\sum_d \log p(w_d) = \sum_d rac{1}{2\beta} |w_d|$



image:https://stats.stackexchange.com/questions/177210/why-is-laplace-prior-producing-sparse-solutions

Laplace prior

another notable choice of prior is the Laplace distribution

minimizing negative log-likelihood
$$\rightarrow -\sum_{d} \log p(w_{d}) = \sum_{d} \frac{1}{2\beta} |w_{d}| = \frac{1}{2\beta} ||w||_{1}$$

L1 norm of w



image:https://stats.stackexchange.com/questions/177210/why-is-laplace-prior-producing-sparse-solutions

Laplace prior

another notable choice of prior is the Laplace distribution



image:https://stats.stackexchange.com/questions/177210/why-is-laplace-prior-producing-sparse-solutions

$L_1 \operatorname{vs} L_2$ regularization

regularization path shows how $\{w_d\}$ change as we change λ



$L_1 \text{ vs } L_2$ regularization

regularization path shows how $\{w_d\}$ change as we change λ Lasso produces sparse weights (many are zero, rather than small)



$L_1 \text{ vs } L_2$ regularization

regularization path shows how $\{w_d\}$ change as we change λ Lasso produces sparse weights (many are zero, rather than small) red-line is the optimal λ from cross-validation



$L_1 \text{ vs } L_2$ regularization

 $\min_w J(w) + \lambda ||w||_p^p$ is equivalent to $\min_w J(w)$ subject to $||w||_p^p \leq \tilde{\lambda}$ for an appropriate choice of $\tilde{\lambda}$

$L_1 \text{ vs } L_2$ regularization

 $\min_w J(w) + \lambda ||w||_p^p$ is equivalent to $\min_w J(w)$ subject to $||w||_p^p \leq \tilde{\lambda}$ for an appropriate choice of $\tilde{\lambda}$ figures below show the constraint and the isocontours of J(w)



$L_1 \text{ vs } L_2$ regularization

 $\min_{w} J(w) + \lambda ||w||_{p}^{p}$ is equivalent to $\min_{w} J(w)$ subject to $||w||_{p}^{p} \leq \tilde{\lambda}$ for an appropriate choice of $\tilde{\lambda}$ figures below show the constraint and the isocontours of J(w)optimal solution with L1-regularization is more likely to have zero components





p-norms with $\,p\geq 1\,{
m are\,\,convex}$ (easier to optimize)



p-norms with $\,p\geq 1\,{
m are\,\,convex}$ (easier to optimize)



p-norms with $\,p\geq 1\,$ are convex (easier to optimize)



p-norms with $\,p\geq 1\,$ are convex (easier to optimize)



penalizes the **number of** non-zero features

 $J(w) + \lambda ||w||_0 = J(w) + \lambda \sum_d \mathbb{I}(w_d
eq 0)$

p-norms with $\,p\geq 1\,$ are convex (easier to optimize)



penalizes the **number of** non-zero features

 $J(w)+\lambda||w||_0=J(w)+\lambda\sum_d\mathbb{I}(w_d
eq 0)$ a penalty of λ for each feature performs feature selection

p-norms with $\,p\geq 1\,$ are convex (easier to optimize)



optimizing this is a difficult *combinatorial problem*:

• search over all 2^D subsets

p-norms with $\,p\geq 1\,$ are convex (easier to optimize)



optimizing this is a difficult *combinatorial problem*:

• search over all 2^D subsets

L1 regularization is a viable alternative to L0 regularization

for L2 loss

for L2 loss

assume a true distribution $\ p(x,y)$

for L2 loss

assume a true distribution $\ p(x,y)$

the regression function is $\ \ f(x) = \mathbb{E}_p[y|x]$

for L2 loss

assume a true distribution $\ p(x,y)$

the regression function is $\ \ f(x) = \mathbb{E}_p[y|x]$

assume that a dataset $\mathcal{D} = \{(x^{(n)},y^{(n)})\}_n$ is sampled from p(x,y)

for L2 loss

assume a true distribution $\ p(x,y)$

the regression function is $\ \ f(x) = \mathbb{E}_p[y|x]$

assume that a dataset $\mathcal{D} = \{(x^{(n)},y^{(n)})\}_n$ is sampled from p(x,y)

let $\hat{f}_{\mathcal{D}}$ be our model based on the dataset

for L2 loss

assume a true distribution p(x, y)the regression function is $f(x) = \mathbb{E}_p[y|x]$ assume that a dataset $\mathcal{D} = \{(x^{(n)}, y^{(n)})\}_n$ is sampled from p(x, y)let $\hat{f}_{\mathcal{D}}$ be our model based on the dataset what we care about is the expected loss (*aka risk*)

 $\mathbb{E}[(\hat{f}_\mathcal{D}(x)-y)^2]$

all blue items are random variables

for L2 loss

$$\mathbb{E}[({\hat f}_{\mathcal D}(x)-y)^2]$$

for L2 loss

$$\mathbb{E}[(\hat{f}_{\mathcal{D}}(x) - oldsymbol{y})^2] egin{array}{c} | \ | \ f(x) + \epsilon \end{array}$$

for L2 loss

$$\mathbb{E}[(\hat{f}_{\mathcal{D}}(x) - y)^2]$$

 $\begin{vmatrix} & | \\ & | \\ f(x) + \epsilon \\ \hat{f}_{\mathcal{D}}(x) + \mathbb{E}_{\mathcal{D}}[\hat{f}_{\mathcal{D}}(x)] - \mathbb{E}_{\mathcal{D}}[\hat{f}_{\mathcal{D}}(x)]$ add and subtract a term

for L2 loss

for L2 loss

what we care about is the expected loss (aka risk)

$$\mathbb{E}[(\hat{f}_{\mathcal{D}}(x) - y)^2] = \mathbb{E}[(\hat{f}_{\mathcal{D}}(x) - \mathbb{E}_{\mathcal{D}}[\hat{f}_{\mathcal{D}}(x)] - y + \mathbb{E}_{\mathcal{D}}[\hat{f}_{\mathcal{D}}(x)])^2]$$

$$\begin{vmatrix} & | \\ & | \\ & f(x) + \epsilon \end{vmatrix}$$

$$\hat{f}_{\mathcal{D}}(x) + \mathbb{E}_{\mathcal{D}}[\hat{f}_{\mathcal{D}}(x)] - \mathbb{E}_{\mathcal{D}}[\hat{f}_{\mathcal{D}}(x)] \text{ add and subtract a term}$$

the remaining terms evaluate to zero (check for yourself!)

 $\mathcal{L} = \mathbb{E}[(\hat{f}_\mathcal{D}(x) - \mathbb{E}_\mathcal{D}[\hat{f}_\mathcal{D}(x)])^2] + \mathbb{E}[(f(x) - \mathbb{E}_\mathcal{D}[\hat{f}_\mathcal{D}(x)])^2] + \mathbb{E}[\epsilon^2]$

for L2 loss

what we care about is the expected loss (aka risk)

$$\mathbb{E}[(\hat{f}_{\mathcal{D}}(x) - y)^2] = \mathbb{E}[(\hat{f}_{\mathcal{D}}(x) - \mathbb{E}_{\mathcal{D}}[\hat{f}_{\mathcal{D}}(x)] - y + \mathbb{E}_{\mathcal{D}}[\hat{f}_{\mathcal{D}}(x)])^2]$$

$$\begin{vmatrix} & | \\ & | \\ & f(x) + \epsilon \\ & \hat{f}_{\mathcal{D}}(x) + \mathbb{E}_{\mathcal{D}}[\hat{f}_{\mathcal{D}}(x)] - \mathbb{E}_{\mathcal{D}}[\hat{f}_{\mathcal{D}}(x)] \text{ add and subtract a term}$$

the remaining terms evaluate to zero (check for yourself!)

 $= \mathbb{E}[(\hat{f}_{\mathcal{D}}(x) - \mathbb{E}_{\mathcal{D}}[\hat{f}_{\mathcal{D}}(x)])^{2}] + \mathbb{E}[(f(x) - \mathbb{E}_{\mathcal{D}}[\hat{f}_{\mathcal{D}}(x)])^{2}] + \mathbb{E}[\epsilon^{2}]$ variance
bias
unavoidable
noise error

the expected loss is decomposed to:

image: P. Domingos' posted article

the expected loss is decomposed to:

 $\mathbb{E}[(f(x) - \mathbb{E}_{\mathcal{D}}[\hat{f}_{\mathcal{D}}(x)])^2]$

bias: how average over all datasets differs from the regression function

the expected loss is decomposed to:

 $\mathbb{E}[(f(x) - \mathbb{E}_{\mathcal{D}}[\hat{f}_{\mathcal{D}}(x)])^2]$

bias: how average over all datasets differs from the regression function

 $\mathbb{E}[(\hat{f}_{\mathcal{D}}(x) - \mathbb{E}_{\mathcal{D}}[\hat{f}_{\mathcal{D}}(x)])^2]$ **variance:** how change of dataset affects the prediction

the expected loss is decomposed to:

 $\mathbb{E}[(f(x) - \mathbb{E}_{\mathcal{D}}[\hat{f}_{\mathcal{D}}(x)])^2]$

bias: how average over all datasets differs from the regression function

 $\mathbb{E}[(\hat{f}_{\mathcal{D}}(x) - \mathbb{E}_{\mathcal{D}}[\hat{f}_{\mathcal{D}}(x)])^2]$ **variance:** how change of dataset affects the prediction

 $\mathbb{E}[\epsilon^2]$

noise error: the error even if we used the true model f(x)

the expected loss is decomposed to:

 $\mathbb{E}[(f(x) - \mathbb{E}_{\mathcal{D}}[\hat{f}_{\mathcal{D}}(x)])^2]$

bias: how average over all datasets differs from the regression function

 $\mathbb{E}[(\hat{f}_{\mathcal{D}}(x) - \mathbb{E}_{\mathcal{D}}[\hat{f}_{\mathcal{D}}(x)])^2]$ **variance:** how change of dataset affects the prediction

 $\mathbb{E}[\epsilon^2]$

noise error: the error even if we used the true model f(x)



the expected loss is decomposed to:

 $\mathbb{E}[(f(x) - \mathbb{E}_{\mathcal{D}}[\hat{f}_{\mathcal{D}}(x)])^2]$

bias: how average over all datasets differs from the regression function

 $\mathbb{E}[(\hat{f}_{\mathcal{D}}(x) - \mathbb{E}_{\mathcal{D}}[\hat{f}_{\mathcal{D}}(x)])^2]$ variance: how change of dataset affects the prediction

 $\mathbb{E}[\epsilon^2]$

noise error: the error even if we used the true model f(x)

different models vary in their trade off between error due to bias and variance

- simple models: often more biased
- complex models: often have more variance



Example: bias vs. variance








variance is the average difference (in squared L2 norm) between these curves and their average

bias is the difference (in L2 norm) between two curves





using larger regularization penalty: higher bias - lower variance





side note

the average fit is very good, despite high variance **model averaging:** uses "average" prediction of expressive models to prevent overfitting

using larger regularization penalty: higher bias - lower variance









Big picture!



Big picture!

high variance in more complex models means that test and training error can be very different



Big picture!

high variance in more complex models means that test and training error can be very different high bias in simplistic models means that training error can be high



Model selection

how to pick the model with lowest expected loss / test error?

USE a validation set (and a separate test set for final assessment)

 Train
 Validation
 Test

 use for model selection
 use for final model assessment

regularization bound the test error by bounding

- training error
- model complexity

Model selection

how to pick the model with lowest expected loss / test error?

USE a validation set (and a separate test set for final assessment)

 Train
 Validation
 Test

 use for model selection
 use for final model assessment

regularization bound the test error by bounding

- training error
- model complexity

in the end we may have to use a validation set to find the right amount of regularization

getting a more reliable estimate of test error using validation set

K-fold cross validation(CV)

- randomly partition the data into K folds
- use K-1 for training, and 1 for validation
- report average/std of the validation error over all folds



getting a more reliable estimate of test error using validation set

K-fold cross validation(CV)

- randomly partition the data into K folds
- use K-1 for training, and 1 for validation
- report average/std of the validation error over all folds



leave-one-out CV:extreme case of k=N

getting a more reliable estimate of test error using validation set

K-fold cross validation(CV)

- randomly partition the data into k *folds*
- use k-1 for training, and 1 for validation
- report average/std of the validation error over all folds



getting a more reliable estimate of test error using validation set

K-fold cross validation(CV)

- randomly partition the data into k *folds*
- use k-1 for training, and 1 for validation
- report average/std of the validation error over all folds



once the hyper-parameters are selected, we can use the whole set for training use test set for the **final** assessment

evaluation metric can be different from the optimization objective confusion matrix is a CxC table that compares truth-vs-prediction

for **binary classification**:

	Truth		Σ
Result	TP	FP	RP
	FN	TN	RN
Σ	Р	Ν	

evaluation metric can be different from the optimization objective **confusion matrix** is a CxC table that compares truth-vs-prediction

for **binary classification**:

	Truth		Σ
Result	TP	FP	RP
	FN	TN	RN
Σ	Р	Ν	

some **evaluation metrics** (based on the confusion table) $Accuracy = \frac{TP+TN}{P+N}$ $Error rate = \frac{FP+FN}{P+N}$ $Precision = \frac{TP}{RP}$ $Recall = \frac{TP}{P}$ $F_1score = 2\frac{Precision \times Recall}{Precision + Recall}$

evaluation metric can be different from the optimization objective confusion matrix is a CxC table that compares truth-vs-prediction

for **binary classification**:

	Truth		Σ
Result	TP	FP	RP
	FN	TN	RN
Σ	Р	Ν	

some **evaluation metrics** (based on the confusion table)

$$Accuracy = rac{TP+TN}{P+N}$$

 $Error \ rate = rac{FP+FN}{P+N}$
 $Precision = rac{TP}{RP}$
 $Recall = rac{TP}{P}$

 $F_1 score = 2 rac{Precision imes Recall}{Precision + Recall}$

type I vs type II error



threshold

if we produce class score (probability) we can trade-off between type I & type II error





goal: evaluate class scores/probabilities (independent of choice of threshold)

if we produce class score (probability) we can trade-off between type I & type II error



goal: evaluate class scores/probabilities (independent of choice of threshold)

Receiver Operating Characteristic ROC curve

TPR = TP/P (**recall**, sensitivity) **FPR** = FP/N (**fallout**, false alarm)



Summary

- complex models can have very different training and test error (generalization gap)
- regularization bounds this gap by penalizing model complexity
 - L1 & L2 regularization
 - probabilistic interpretation: different priors on weights
 - L1 produces sparse solutions (useful for feature selection)

Summary

- complex models can have very different training and test error (generalization gap)
- regularization bounds this gap by penalizing model complexity
 - L1 & L2 regularization
 - probabilistic interpretation: different priors on weights
 - L1 produces sparse solutions (useful for feature selection)
- bias-variance trade off:
 - formalizes the relation between
 - training error (bias)
 - complexity (variance) and
 - and the test error (bias + variance)
 - not so elegant beyond L2 loss

Summary

- complex models can have very different training and test error (generalization gap)
- regularization bounds this gap by penalizing model complexity
 - L1 & L2 regularization
 - probabilistic interpretation: different priors on weights
 - L1 produces sparse solutions (useful for feature selection)
- bias-variance trade off:
 - formalizes the relation between
 - training error (bias)
 - complexity (variance) and
 - and the test error (bias + variance)
 - not so elegant beyond L2 loss
- (cross) validation for model selection