Learning probability distributions

Abbas Mehrabian

McGill University IVADO Postdoctoral Fellow

29 November 2018

Co-authors: Hassan Ashtiani, Shai Ben-David, Luc Devroye, Nick Harvey, Chris Liaw, Yaniv Plan, and Tommy Reddad

An example of distribution learning Generating random faces for computer games

- $\checkmark\,$ Training data consists of actual faces.
- A probability density function $\mathbf{P}: \mathbb{R}^d \to \mathbb{R}$ is learned from the data.
- $\checkmark\,$ New random faces are generated using the learned distribution.

An example of distribution learning Generating random faces for computer games

- $\checkmark\,$ Training data consists of actual faces.
- A probability density function $\mathbf{P}: \mathbb{R}^d \to \mathbb{R}$ is learned from the data.
- $\checkmark\,$ New random faces are generated using the learned distribution.

A popular approach: generative adversarial networks (GANs), based on deep neural networks.

Distribution learning in action



Top: generated images using generative adversarial networks Bottom: a small part of the training data

> Picture from Karras, Aila, Laine, and Lehtinen (NVIDIA and Aalto University), October 2017

also known as density estimation

Given an i.i.d. sample generated from an unknown target distribution $\hat{\mathbf{P}}$, output a distribution $\hat{\mathbf{P}}$ that is close to \mathbf{P} .

also known as density estimation

Given an i.i.d. sample generated from an unknown target distribution $\hat{\mathbf{P}}$, output a distribution $\hat{\mathbf{P}}$ that is close to \mathbf{P} .

- $\checkmark~$ We assume ${\bf P}$ belongs to some known class ${\cal F}$ of distributions.
- $\checkmark\,$ We would like our algorithm to use as a small sample as possible.
- $\checkmark \ \text{Closeness is measured by the total variation distance:} \\ \mathrm{TV}(\mathbf{P},\widehat{\mathbf{P}}) \coloneqq \sup_E |\mathbf{P}(E) \widehat{\mathbf{P}}(E)| = \frac{1}{2} \int |p(x) \widehat{p}(x)| \, \mathrm{d}x$

Our set up

Given an i.i.d. sample generated from an unknown target distribution \mathbf{P} from a known class \mathcal{F} , output some $\widehat{\mathbf{P}}$ that is close to \mathbf{P} .

What is the smallest number of samples needed to guarantee $TV(\widehat{\mathbf{P}}, \mathbf{P}) \leq \varepsilon$ with probability 99%? $m_{\mathcal{F}}(\varepsilon)$.

Our set up

Given an i.i.d. sample generated from an unknown target distribution \mathbf{P} from a known class \mathcal{F} , output some $\widehat{\mathbf{P}}$ that is close to \mathbf{P} .

What is the smallest number of samples needed to guarantee $TV(\widehat{\mathbf{P}}, \mathbf{P}) \leq \varepsilon$ with probability 99%? $m_{\mathcal{F}}(\varepsilon)$.

Main problem

prove bounds for $m_{\mathcal{F}}(\varepsilon)$ for various classes \mathcal{F} .

Our set up

Given an i.i.d. sample generated from an unknown target distribution \mathbf{P} from a known class \mathcal{F} , output some $\widehat{\mathbf{P}}$ that is close to \mathbf{P} .

What is the smallest number of samples needed to guarantee $TV(\widehat{\mathbf{P}}, \mathbf{P}) \leq \varepsilon$ with probability 99%? $m_{\mathcal{F}}(\varepsilon)$.

Main problem

prove bounds for $m_{\mathcal{F}}(\varepsilon)$ for various classes \mathcal{F} .

Often in statistics the problem is stated differently: given n samples from \mathbf{P} , how small can you make $\mathbb{E} \operatorname{TV}(\widehat{\mathbf{P}}, \mathbf{P})$? The answer is called the minimax risk of \mathcal{F} .

A heuristic

 $arepsilon^2 m_{\mathcal{F}}(arepsilon) symp {
m number}$ of free parameters in $\mathcal F$ in 'natural representation'

Example

- \checkmark $\mathcal{F}=$ Bernoulli distributions: $m_{\mathcal{F}}(\varepsilon) symp 1/\varepsilon^2$
- $\checkmark \mathcal{F} = ext{Gaussian distributions:} \ m_{\mathcal{F}}(\varepsilon) symp 2/arepsilon^2$
- $\checkmark \ \mathcal{F} = d$ -dimensional Gaussians: $m_{\mathcal{F}}(\varepsilon) \leq C d^2 / \varepsilon^2$
- \checkmark Finite $\mathcal{F}: \ m_{\mathcal{F}}(\varepsilon) \leq C \log |\mathcal{F}|/\varepsilon^2$ Devroye-Lugosi'01

A heuristic

 $arepsilon^2 m_{\mathcal{F}}(arepsilon) symp {
m number}$ of free parameters in $\mathcal F$ in 'natural representation'

Example

- $\checkmark \ \mathcal{F} = ext{Bernoulli distributions:} \ m_{\mathcal{F}}(\varepsilon) symp 1/arepsilon^2$
- $\checkmark \mathcal{F} = ext{Gaussian distributions:} \ m_{\mathcal{F}}(\varepsilon) symp 2/\varepsilon^2$
- $\checkmark~{\cal F}=d$ -dimensional Gaussians: $m_{\cal F}(arepsilon)\leq Cd^2/arepsilon^2$
- \checkmark Finite \mathcal{F} : $m_{\mathcal{F}}(\varepsilon) \leq C \log |\mathcal{F}|/\varepsilon^2$ Devroye-Lugosi'01

Main contribution: this heuristic also works for two more complicated classes: mixtures of multidimensional Gaussians, and the Ising model.

Mixtures of Gaussians

Mixtures of Gaussians

A mixture of k Gaussians in d dimensions has density $\sum_{i=1}^{k} w_i \mathcal{N}(\mu_i, \Sigma_i)(x)$, where $w_i \ge 0$ and $\sum w_i = 1$. $\mathcal{N}(\mu, \Sigma)(x) = \text{density of a Gaussian with mean } \mu \in \mathbb{R}^d$ and covariance matrix $\Sigma \in \mathbb{R}^{d \times d}$





Main result

mixtures of Gaussians

Theorem (Ashtiani, Ben-David, Harvey, Liaw, M, Plan'18)

Let $\mathcal{F}_{k,d} = mixtures$ of k Gaussians in ddimensions. Then, $m_{\mathcal{F}_{k,d}}(\varepsilon) = kd^2/\varepsilon^2$ up to polylogarithmic factors.

Main result

mixtures of Gaussians

Theorem (Ashtiani, Ben-David, Harvey, Liaw, M, Plan'18)

Let $\mathcal{F}_{k,d} = mixtures$ of k Gaussians in ddimensions. Then, $m_{\mathcal{F}_{k,d}}(\varepsilon) = kd^2/\varepsilon^2$ up to polylogarithmic factors.

Any density in $\mathcal{F}_{k,d}$ has form $\sum_{i=1}^{k} w_i \mathcal{N}(\mu_i, \Sigma_i)$, and Σ_i is $d \times d$, so has $\Theta(kd^2)$ parameters.

Definition

 \mathcal{F} admits τ -compression, if for any $\mathbf{P} \in \mathcal{F}$, there exist τ data points from which \mathbf{P} can be reconstructed.

Definition

 \mathcal{F} admits τ -compression, if for any $\mathbf{P} \in \mathcal{F}$, there exist τ data points from which \mathbf{P} can be reconstructed.



Definition

 \mathcal{F} admits τ -compression, if for any $\mathbf{P} \in \mathcal{F}$, you can find τ data points from which \mathbf{P} can be reconstructed.



Definition

 \mathcal{F} admits τ -compression, if for any $\mathbf{P} \in \mathcal{F}$, you can find τ data points from which \mathbf{P} can be reconstructed.



Definition

 \mathcal{F} admits τ -compression, if for any $\mathbf{P} \in \mathcal{F}$, you can find τ data points from which \mathbf{P} can be reconstructed.



Definition

 \mathcal{F} admits τ -compression, if for any $\mathbf{P} \in \mathcal{F}$, you can find τ data points from which \mathbf{P} can be reconstructed.

$$\begin{array}{c} \widehat{\mu} = \frac{x_1 + x_2}{2} \\ \widehat{\sigma} = \frac{|x_1 - x_2|}{2} \end{array} \\ \hline x_1 \qquad \qquad x_2 \end{array}$$

1. Compressing *d*-dimensional Gaussians

d-dimensional Gaussians admit $O(d^2 \log d)$ -compression.

2. Compressing mixtures

If \mathcal{F} admits τ -compression, then k-mix (\mathcal{F}) admits $(k\tau + k \log k)$ -compression.

3. Compression implies learnability

If \mathcal{F} admits τ -compression, then $m_{\mathcal{F}}(\varepsilon) = O(\tau \log \tau / \varepsilon^2)$.

Theorem (Ashtiani, Ben-David, Harvey, Liaw, M, Plan'18)

Let $\mathcal{F}_{k,d} = mixtures$ of k Gaussians in d dimensions. Then, $m_{\mathcal{F}_{k,d}}(\varepsilon) \leq (kd^2/\varepsilon^2) \times \operatorname{polylog}(kd/\varepsilon).$

1. Compressing *d*-dimensional Gaussians

d-dimensional Gaussians admit $O(d^2 \log d)$ -compression.

Suppose
$$d = 2$$
, consider $\mathbf{P} = \mathcal{N}(\mu, \Sigma) = \mathcal{N}(\mu, v_1 v_1^{\mathsf{T}} + v_2 v_2^{\mathsf{T}})$.

1. Compressing *d*-dimensional Gaussians

d-dimensional Gaussians admit $O(d^2 \log d)$ -compression.



1. Compressing *d*-dimensional Gaussians

d-dimensional Gaussians admit $O(d^2 \log d)$ -compression.



1. Compressing *d*-dimensional Gaussians

d-dimensional Gaussians admit $O(d^2 \log d)$ -compression.



1. Compressing d-dimensional Gaussians

d-dimensional Gaussians admit $O(d^2 \log d)$ -compression.



1. Compressing *d*-dimensional Gaussians

d-dimensional Gaussians admit $O(d^2 \log d)$ -compression.



1. Compressing d-dimensional Gaussians

d-dimensional Gaussians admit $O(d^2 \log d)$ -compression.



1. Compressing d-dimensional Gaussians

d-dimensional Gaussians admit $O(d^2 \log d)$ -compression.



1. Compressing *d*-dimensional Gaussians

d-dimensional Gaussians admit $O(d^2 \log d)$ -compression.



1. Compressing *d*-dimensional Gaussians

d-dimensional Gaussians admit $O(d^2 \log d)$ -compression.

Suppose d = 2, consider $\mathbf{P} = \mathcal{N}(\mu, \Sigma) = \mathcal{N}(\mu, v_1 v_1^{\mathsf{T}} + v_2 v_2^{\mathsf{T}})$.



For d > 2, use $d \log d$ data points to 'encode' the mean, and $d \log d$ data points for each eigenvector.

2. Compressing mixtures

If \mathcal{F} admits τ -compression, then k-mix (\mathcal{F}) admits $(k\tau + k \log k)$ -compression.

2. Compressing mixtures

If \mathcal{F} admits τ -compression, then k-mix (\mathcal{F}) admits $(k\tau + k \log k)$ -compression.

. . .

2. Compressing mixtures

If \mathcal{F} admits τ -compression, then k-mix (\mathcal{F}) admits $(k\tau + k \log k)$ -compression.



2. Compressing mixtures

If \mathcal{F} admits τ -compression, then k-mix (\mathcal{F}) admits $(k\tau + k \log k)$ -compression.



2. Compressing mixtures

If \mathcal{F} admits τ -compression, then k-mix (\mathcal{F}) admits $(k\tau + k \log k)$ -compression.



2. Compressing mixtures

If \mathcal{F} admits τ -compression, then k-mix (\mathcal{F}) admits $(k\tau + k \log k)$ -compression.

Let
$$\widehat{\mathbf{P}} = \frac{1}{2}\widehat{P_1} + \frac{1}{2}\widehat{P_2} + \frac{1}{2}\widehat{P_3}$$

3. Compression implies learnability

If \mathcal{F} admits τ -compression, then $m_{\mathcal{F}}(\varepsilon) = O(\tau \log \tau / \varepsilon^2)$.

First, generate a sample of size $m = \text{poly}(\tau)$. Try to reconstruct the distribution by considering all $\binom{m}{\tau}$ subsets of size τ (we know one of them is correct).

3. Compression implies learnability

If \mathcal{F} admits τ -compression, then $m_{\mathcal{F}}(\varepsilon) = O(\tau \log \tau / \varepsilon^2)$.

First, generate a sample of size $m = \text{poly}(\tau)$. Try to reconstruct the distribution by considering all $\binom{m}{\tau}$ subsets of size τ (we know one of them is correct).

Theorem (Devroye and Lugosi'01)

Given a finite set C of candidates, given $\log(|C|)/\epsilon^2$ additional samples from the target distribution, we can find the candidate that is closest to the target.

In our case, $|C| = \binom{m}{\tau} \le m^{\tau}$, hence total sample complexity $< \tau \log(m)/\varepsilon^2$.

1. Compressing *d*-dimensional Gaussians

d-dimensional Gaussians admit $O(d^2 \log d)$ -compression.

2. Compressing mixtures

If \mathcal{F} admits τ -compression, then k-mix (\mathcal{F}) admits $(k\tau + k \log k)$ -compression.

3. Compression implies learnability

If \mathcal{F} admits τ -compression, then $m_{\mathcal{F}}(\varepsilon) = O(\tau \log \tau / \varepsilon^2)$.

Theorem (Ashtiani, Ben-David, Harvey, Liaw, M, Plan'18)

Let $\mathcal{F}_{k,d} = mixtures$ of k Gaussians in d dimensions. Then, $m_{\mathcal{F}_{k,d}}(\varepsilon) = \widetilde{O}(kd^2/\varepsilon^2).$

Main lemma

Let $\mathcal{F}_{1,d} = d$ -dimensional Gaussians. Then, $m_{\mathcal{F}_{1,d}}(\varepsilon) = \widetilde{\Omega}(d^2/\varepsilon^2).$

Main lemma

Let
$$\mathcal{F}_{1,d} = d$$
-dimensional Gaussians. Then,
 $m_{\mathcal{F}_{1,d}}(\varepsilon) = \widetilde{\Omega}(d^2/\varepsilon^2).$

Fano's lemma

Suppose there exist $f_1,\ldots,f_M\in\mathcal{F}$ with

 $\operatorname{KL}(f_i \parallel f_j) = O(\varepsilon^2) ext{ and } \operatorname{TV}(f_i, f_j) = \Omega(\varepsilon) \qquad \forall i \neq j \in [M].$

Then $m_{\mathcal{F}}(\varepsilon) = \Omega(\log M / \varepsilon^2)$.

$$ext{KL}(f_1 \parallel f_2) \coloneqq \int f_1(x) \log rac{f_1(x)}{f_2(x)} \mathrm{d}x$$

Main lemma

Let $\mathcal{F}_{1,d} = d$ -dimensional Gaussians. Then, $m_{\mathcal{F}_{1,d}}(\varepsilon) = \widetilde{\Omega}(d^2/\varepsilon^2).$

Fano's lemma

Suppose there exist $f_1,\ldots,f_M\in\mathcal{F}$ with

 $\operatorname{KL}(f_i \parallel f_j) = O(\varepsilon^2) \text{ and } \operatorname{TV}(f_i, f_j) = \Omega(\varepsilon) \qquad \forall i \neq j \in [M].$

Then $m_{\mathcal{F}}(\varepsilon) = \Omega(\log M / \varepsilon^2)$.

$$ext{KL}(f_1 \parallel f_2) \coloneqq \int f_1(x) \log rac{f_1(x)}{f_2(x)} \mathrm{d}x$$

To apply this lemma, we need to build 2^{d^2} Gaussian distributions, with pairwise KL-divergence $\leq \varepsilon^2$, pairwise TV distance $\geq \varepsilon$.

Need to build 2^{d^2} Gaussian distributions with pairwise KL-divergence $\leq \varepsilon^2$ and pairwise TV distance $\geq \varepsilon$. We will use zero-mean Gaussians, so just need to specify the covariance matrices. Need to build 2^{d^2} Gaussian distributions with pairwise KL-divergence $\leq \varepsilon^2$ and pairwise TV distance $\geq \varepsilon$. We will use zero-mean Gaussians, so just need to specify the covariance matrices.

First construction (geometric). Repeat 2^{d^2} times: start with an identity covariance matrix, then choose a random subspace of dimension d/9 and slightly increase the eigenvalues corresponding to this eigenspace: $\Sigma = I + \frac{\varepsilon}{\sqrt{d}} UU^{\mathsf{T}}$, with $U \in \mathbb{R}^{d \times d/9}$ orthonormal.

Then prove that with large probability, any two of these have TV distance $\geq \epsilon.$

Need to build 2^{d^2} Gaussian distributions with pairwise KL-divergence $\leq \epsilon^2$ and pairwise TV distance $\geq \epsilon$. We will use zero-mean Gaussians, so just need to specify the covariance matrices.

Second construction (combinatorial). For d = 3, consider the following inverse covariance matrices:

$$\begin{pmatrix} 0 & -\delta & -\delta \\ -\delta & 0 & -\delta \\ -\delta & -\delta & 0 \end{pmatrix}, \begin{pmatrix} 0 & \delta & \delta \\ \delta & 0 & -\delta \\ \delta & -\delta & 0 \end{pmatrix}, \begin{pmatrix} 0 & \delta & -\delta \\ \delta & 0 & \delta \\ -\delta & \delta & 0 \end{pmatrix}, \begin{pmatrix} 0 & -\delta & \delta \\ -\delta & 0 & \delta \\ \delta & \delta & 0 \end{pmatrix}$$

For general d, build $2^{d^2/10}$ inverse covariance matrices so that any two of them are different in at least $d^2/3$ coordinates.

Main lemma

Let $\mathcal{F}_{1,d} = d$ -dimensional Gaussians. Then, $m_{\mathcal{F}_{1,d}}(\varepsilon) = \widetilde{\Omega}(d^2/\varepsilon^2).$

It is easy to lift this to the class of mixtures, proving $m_{\mathcal{F}_{k,d}}(\varepsilon) = \widetilde{\Omega}(kd^2/\varepsilon^2).$

Main lemma

Let $\mathcal{F}_{1,d} = d$ -dimensional Gaussians. Then, $m_{\mathcal{F}_{1,d}}(\varepsilon) = \widetilde{\Omega}(d^2/\varepsilon^2).$

It is easy to lift this to the class of mixtures, proving $m_{\mathcal{F}_{k,d}}(\varepsilon) = \widetilde{\Omega}(kd^2/\varepsilon^2).$

Theorem (Ashtiani, Ben-David, Harvey, Liaw, M, Plan'18)

Let $\mathcal{F}_{k,d}$ = mixtures of k Gaussians in d dimensions. Then, $m_{\mathcal{F}_{k,d}}(\varepsilon) = kd^2/\varepsilon^2$ up to polylogarithmic factors. The Ising model

The Ising model

Definition

For a graph G on d vertices, and edge weights $\{w_{i,j}\}_{ij \in E(G)}$, the Ising model with parameters $\{w_{i,j}\}_{ij \in E(G)}$ is supported on $\{-1, +1\}^d$ and has probability mass function

$$p_{\mathbf{w}}(x_1,\ldots,x_d) \propto \exp\left(\sum_{ij \,\in\, E(\,G)} w_{i,j} x_i x_j
ight)$$

Number of parameters = |E(G)|.

The Ising model

Definition

For a graph G on d vertices, and edge weights $\{w_{i,j}\}_{ij \in E(G)}$, the Ising model with parameters $\{w_{i,j}\}_{ij \in E(G)}$ is supported on $\{-1,+1\}^d$ and has probability mass function

$$p_{\mathbf{w}}(x_1,\ldots,x_d) \propto \exp\left(\sum_{ij \,\in\, E(\,G)} w_{i,j} x_i x_j
ight)$$

Number of parameters = |E(G)|.

Theorem (Devroye, M, Reddad'18)

Let $\mathcal{I}_G = Ising models$ on G. Then, $m_{\mathcal{I}_G}(\varepsilon) \asymp |E(G)|/\varepsilon^2$.

Lower bound proof uses Fano's inequality again. Need to build $2^{|E(G)|}$ Ising models with pairwise KL-divergence $\leq \epsilon^2$ and pairwise TV distance $\geq \epsilon$. Lower bound proof uses Fano's inequality again. Need to build $2^{|E(G)|}$ Ising models with pairwise KL-divergence $\leq \epsilon^2$ and pairwise TV distance $\geq \epsilon$.

For d = 3 and G the complete graph, consider the following weight matrices W:

$$\begin{pmatrix} 0 & -\delta & -\delta \\ -\delta & 0 & -\delta \\ -\delta & -\delta & 0 \end{pmatrix}, \begin{pmatrix} 0 & \delta & \delta \\ \delta & 0 & -\delta \\ \delta & -\delta & 0 \end{pmatrix}, \begin{pmatrix} 0 & \delta & -\delta \\ \delta & 0 & \delta \\ -\delta & \delta & 0 \end{pmatrix}, \begin{pmatrix} 0 & -\delta & \delta \\ -\delta & 0 & \delta \\ \delta & \delta & 0 \end{pmatrix}$$

For a general interaction graph G, build $2^{|E(G)|/5}$ weight matrices so that any two of them are different in at least |E(G)|/6 coordinates.

Proof of upper bound

For class \mathcal{F} of densities defined over X, consider the Yatracos set system:

 $A_{\mathcal{F}} \coloneqq \{S \subseteq X : \exists p_1, p_2 \in \mathcal{F} \text{ s. t. } S = \{x \in X : p_1(x) > p_2(x)\}\}$

Proof of upper bound

For class \mathcal{F} of densities defined over X, consider the Yatracos set system:

$$A_\mathcal{F}\coloneqq \{S\subseteq X: \exists p_1,p_2\in \mathcal{F} ext{ s. t. } S=\{x\in X: p_1(x)>p_2(x)\}\}$$

Devroye and Lugosi'01 proved $m_{\mathcal{F}}(\varepsilon) \leq C \cdot \text{VC-dim}(A_{\mathcal{F}})/\varepsilon^2$.

Proof of upper bound

For class \mathcal{F} of densities defined over X, consider the Yatracos set system:

$$A_\mathcal{F}\coloneqq \{S\subseteq X: \exists p_1,p_2\in \mathcal{F} ext{ s. t. } S=\{x\in X: p_1(x)>p_2(x)\}\}$$

Devroye and Lugosi'01 proved $m_{\mathcal{F}}(\varepsilon) \leq C \cdot \text{VC-dim}(A_{\mathcal{F}})/\varepsilon^2$.

If \mathcal{F} is the class of Ising models on G, standard techniques give $\operatorname{VC-dim}(A_{\mathcal{F}}) \leq |E(G)| + 1$, whence $m_{\mathcal{F}}(\varepsilon) \leq C(|E(G)| + 1)/\varepsilon^2$.

Theorem (Devroye, M, Reddad'18)

Let $\mathcal{I}_G = Ising models$ on G. Then, $m_{\mathcal{I}_G}(\varepsilon) \asymp |E(G)|/\varepsilon^2$.

Recap

 $arepsilon^2 m_{\mathcal{F}}(arepsilon) symp ext{number}$ of free parameters in \mathcal{F} in 'natural representation'

Example

- \checkmark $\mathcal{F}=$ Bernoulli distributions: $m_{\mathcal{F}}(arepsilon) pprox 1/arepsilon^2$
- $\checkmark \mathcal{F} = ext{Gaussian distributions:} \ m_{\mathcal{F}}(\varepsilon) symp 1/arepsilon^2$
- $\checkmark~{\cal F}=d$ -dimensional Gaussian distributions: $m_{{\cal F}}(\epsilon) symp d^2/\epsilon^2$
- \checkmark Finite \mathcal{F} : $m_{\mathcal{F}}(\varepsilon) \leq 9 \log |\mathcal{F}|/\varepsilon^2$ Devroye-Lugosi'01
- $\checkmark \mathcal{F}_{k,d} = \text{mixture of } k \text{ Gaussians in } d \text{ dimensions:}$ $m_{\mathcal{F}_{k,d}}(\varepsilon) = \widetilde{\Theta}(kd^2/\varepsilon^2).$
- $\checkmark \ \mathcal{I}_G = \text{Ising models on } G: \ m_{\mathcal{I}_G}(\varepsilon) \asymp |E(G)| / \varepsilon^2.$

Questions

 $arepsilon^2 m_{\mathcal{F}}(arepsilon) symp {
m number}$ of free parameters in $\mathcal F$ in 'natural representation'

- 1. Does the heuristic works for other classes? For example, other exponential families, graphical models, distributions generated by neural networks?
- 2. $\varepsilon^2 m_{\mathcal{F}}(\varepsilon) \leq \text{smallest compression size of } \mathcal{F}$. Is the converse true?
- Can we use ε²m_F(ε) as a natural definition of 'dimension' for class F ? Are there connections with other dimensions?
- 4. What about computational complexity?