Managing the currency risk of a futures portfolio

Alexandre BEAULNE, HEC Montréal

supervisors: Bruno RÉMILLARD, HEC Montréal Pierre LAROCHE, National Bank of Canada

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Problem: Global investors often need to post collateral in multiple currencies, while their performance is measured in one currency, creating exchange rate risk.

Two competing incentives:

- ► Keep posted collateral *low* to minimize exchange rate risk
- Keep posted collateral high to minimize margin calls

What collateral levels in each different currencies optimally balance these two opposing forces?

Optimal collateral

Similar problems:

- Equity portfolio hedging -Minimizing currency risk while minimizing insurance costs
- Transaction costs -Maximizing risk/return while minimizing transaction costs

- Inventory management -Maximizing sales while minimizing shipping costs
- Staff dispatch -Minimizing travel time while minimizing expenses

A good solution needs to properly forecast the underlying prices and exchange rates, accounting for the higher moments and comoments of their respective time-series.

What we did:

- Select a few candidate models for the dynamics of the underlyings' prices and exchange rates
- Assess and compare their goodness-of-fit
- Optimize the "portfolio" of posted collateral based on the chosen model

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Model

We chose a copula-based multivariate GARCH framework as advocated in Xiaohong and Yanqin [2006], Patton [2006] and Rémillard [2010]:

$$X_{i,t} = \mu_t(\boldsymbol{\theta}_i) + h_t(\boldsymbol{\theta}_i)^{1/2} \epsilon_{i,t}$$
(1)

where i = 1, ..., D and innovations $\epsilon_{1,t}, ..., \epsilon_{D,t}$ are *i.i.d.* with continous multivariate distribution function

$$K(x_1,\ldots,x_D)=C_{\theta}(F_1(x_1),\ldots,F_D(x_D))$$
(2)

where the F_i are the cumulative distribution functions of the marginal distributions X_i and C_{θ} is the copula function with parameter(s) θ .

Model

Two steps:

(i) Find appropriate univariate process for each random variable (e.g. AR(1)-GARCH(1,1), eGARCH, GJR-GARCH, etc) for

$$X_{i,t} = \mu_t(\boldsymbol{\theta}_i) + h_t(\boldsymbol{\theta}_i)^{1/2} \epsilon_{i,t}$$

 (ii) Find appropriate copula to capture the dependence between the standardized residuals (e.g. Gaussian, Student, Clayton, Frank, Gumbel, etc) for

$$C_{\theta}(F_1(x_1),\ldots,F_D(x_D))$$

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How do you choose between the different models and once a model is chosen, how do you know it is statistically correct?

 \Rightarrow parametric bootstrapping

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Goodness-of-fit

 H_0 : Dataset belongs to said distribution H_1 : Dataset does not belong to said distribution

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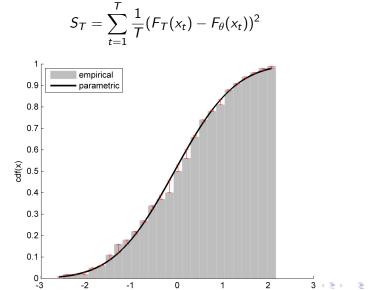
Parametric bootstrapping

General procedure:

- (i) Estimate the parameters of the chosen parametric distribution that best fit the dataset
- (ii) Calculate a distance S_T between the empirical distribution and the parametric distribution (good candidate: Cramèr-von Mises statistic)
- (iii) Generate a large number N of "bootstrapped" samples of the same size as the dataset from the parametric distribution
- (iv) For each of these bootstrapped samples $k = 1, \ldots, N$,
 - (a) Estimate the parameters of the chosen parametric distribution that best fit the bootstrapped sample
 - (b) Calculate a distance $S_T^{(k)}$ between their empirical distribution and the parametric distribution
- (v) The *p*-value for the test is given by the fraction of the $S_T^{(k)}$ bigger than S_T

Cramèr-von Mises statistic

For univariate distributions, the Cramèr-von Mises statistic is given by



Cramèr-von Mises statistic

For copulas, the Cramèr-von Mises statistic is given by

$$S_T = \sum_{t=1}^T \frac{1}{T} (C_T(\hat{u}_{1,t},\ldots,\hat{u}_{D,t}) - C_{\theta}(\hat{u}_{1,t},\ldots,\hat{u}_{D,t}))^2$$

where $\hat{u}_{1,t},\ldots,\hat{u}_{D,t}$ are the normalized ranks

$$\hat{u}_{i,t} = \frac{1}{T-1} \sum_{k=1}^{T} \mathbb{1}(x_{i,t} \geq x_{i,k}),$$

 C_T is the empirical copula

$$C_T(u_{1,t},\ldots,u_{D,t}) = \frac{1}{T-1} \sum_{k=1}^T \mathbb{1}(\hat{u}_{1,t} \ge u_{i,k},\ldots,\hat{u}_{D,t} \ge u_{D,k})$$

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and C_{θ} is the parametric copula chosen.

Unfortunately, C_{θ} do not often have a closed form and numerical approximations are computationally impractical when the number of dimensions gets high. Fortunately an alternative is proposed in Genest et al. [2009] using Rosenblatt's transform:

 $\mathbf{U} \sim \mathcal{C} \ \Leftrightarrow \ \mathcal{T}(\mathbf{U}) \sim \mathcal{C}_{\perp}$

Rosenblatt transform

$$\mathcal{T}(u_1,\ldots,u_D)=(e_1,\ldots,e_D)$$
 given by $e_1=u_1$ and $rac{\delta^{i-1}}{\delta u_1-\delta u_2-1}$ $\mathcal{C}(u_1,\ldots,u_i,1,\ldots,1)$

$$e_{i} = \frac{\frac{\delta u_{1}...\delta u_{i-1}}{\delta u_{1}...\delta u_{i-1}} C(u_{1},...,u_{i},1,...,1)}{\frac{\delta^{i-1}}{\delta u_{1}...\delta u_{i-1}} C(u_{1},...,u_{i-1},1,...,1)}$$
(3)

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[Rosenblatt, 1952].

The recipe to compute the Rosenblatt transform for both meta-elliptical and archimedean copulas can be found in Rémillard et al. [2011].

Parametric bootstrapping - Copula-based Multivariate GARCH model

(i) Estimate the parameters of each univariate marginal process
(ii) Estimate the parameter(s) of the chosen copula on the standardized residuals *e*_t obtained in step (i)

(iii) Compute the normalized ranks $\mathbf{u}_t = u_{1,t}, \dots, u_{D,t}$:

$$u_{i,t} = \frac{1}{T-1} \sum_{k=1}^{T} \mathbb{1}(\epsilon_{i,t} \geq \epsilon_{i,k})$$

Parametric bootstrapping - Copula-based Multivariate GARCH model

(iv) Compute Rosenblatt transforms e_t = e_{1,t},..., e_{D,t}, t = 1,..., T using equation (3)
(v) Compute Cramér-von Mises statistic

$$S_{T} = T \int_{[0,1]^{D}} \{F_{T}(\mathbf{u}) - C_{\perp}(\mathbf{u})\}^{2} d\mathbf{u}$$

= $\frac{T}{3^{D}} - \frac{1}{2^{D-1}} \sum_{t=1}^{T} \prod_{i=1}^{D} (1 - e_{i,t}^{2}) + \frac{1}{T} \sum_{t=1}^{T} \sum_{k=1}^{T} \prod_{i=1}^{D} (1 - \max(e_{i,t}, e_{i,k}))$

Parametric bootstrapping - Copula-based Multivariate GARCH model

- (vi) For some large integer N, repeat the following steps for each k in (1, ..., N):
 - (a) Generate random trajectories of the processes with parameters found in (i) and (ii) of the same length as the original dataset
 - (b) Repeat steps (i) to (v) on trajectories generated in (a) to obtain S^(k)_T.
- (vii) The approximate *p*-value for the test is given by

$$p = \frac{1}{N} \sum_{k=1}^{N} \mathbb{1}(S_T^{(k)} > S_T)$$

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The optimization objective:

Minimize the exchange rate risk on the posted collateral (as measured by the tracking error, Value-at-Risk or Tail Conditional Expectation) subject to a given tolerance on the probability of a margin call

Collateral optimization

In mathematical terms:

$$\min_{\lambda_{1,t},\dots,\lambda_{D,t}} R^{\alpha} \left(\sum_{i=1}^{D} \left(\lambda_{i,t} - \lambda_{i,t}^{*} \right) \times Y_{i,t+1} \right) \tag{4}$$
subject to
$$\lambda_{i,t}^{*} \leq \lambda_{i,t} \leq \infty, \ i = 1,\dots,D$$
and
$$\mathbb{P} \left(\left(\prod_{i=1}^{D} \mathbb{1} \left(\lambda_{i,t} + \mathsf{PnL}_{i,t+1} \geq \lambda_{i,t}^{*} \right) \right) = 0 \right) \leq P_{\mathsf{tol}} \tag{5}$$

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Collateral optimization

where

$$R^{\alpha}(X) = \begin{cases} \mathbb{E}\left[|X|\right] & \text{for expected tracking error} \\ -x^{(\alpha)}(X) & \text{for Value-at-Risk} \\ \mathbb{E}[X|X \le x^{(\alpha)}] & \text{for Tail Conditional Expectation} \end{cases}$$
(6)

and

$$\mathsf{PnL}_{i,t+1} = \sum_{j=1}^{n_{i,t}} {}_{i}\omega_{j,t} \times {}_{i}W_{j,t+1},$$

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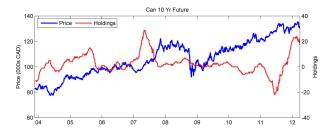
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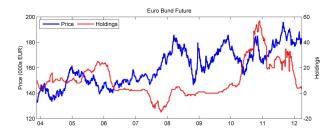
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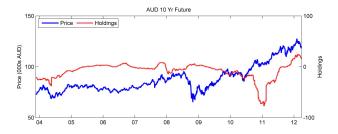
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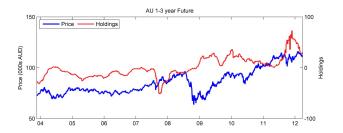
The data set consists of daily holdings of five futures contracts denominated in four non-USD currencies from november 2003 to march 2012 for a total of 1941 observations.

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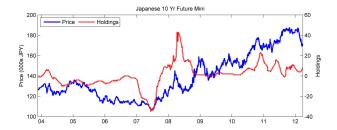




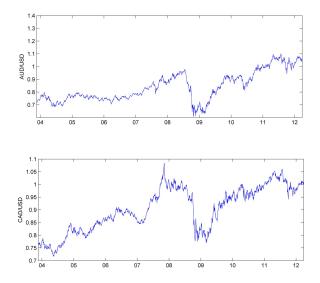




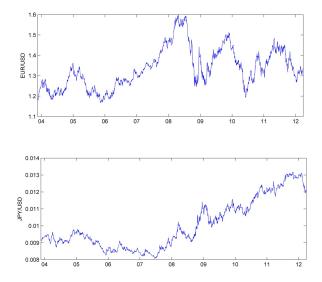
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	Japanese 10 Yr Future Min	Can 10 Yr Future	Euro Bund Future	AUD 10 Yr Future	AU 1-3 year Future	AUD-USD	CAD-USD	EUR-USD	asu-yqu	
AR(1)-GARCH(1,1), gaussian innovations	0.00	0.02	0.00	0.00	0.00	0.00	0.00	0.00	0.00	
AR(2)-GARCH(2,2), gaussian innovations	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.01	0.00	
AR(1)-GARCH(1,1), student innovations	0.56	0.69	0.37	0.58	0.50	0.45	0.60	0.54	0.52	

Table : p-values from the goodness-of-fit tests on marginal processes

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	p values
MV Gaussian	0.00
AR(1)-GARCH(1,1) & gaussian copula	0.01
AR(1)-GARCH(1,1) & student copula	0.12
AR(1)-GARCH(1,1) & Clayton copula	0.00
AR(1)-GARCH(1,1) & Frank copula	0.00
AR(1)-GARCH(1,1) & Gumbel copula	0.00

Table : *p*-values from the goodness-of-fit tests on copula-based MV GARCH models

- Two alternative strategies:
 - (i) Naive: Always post as collateral 2x the minimum margins requirements

- (ii) Model the nine time series with a multivariate Gaussian
- ▶ 500 days buffer left at beginning of sample for calibration
- Daily recalibration
- GOF tests run every year
- $P_{tol} = 0.05, \ \alpha = 0.05$

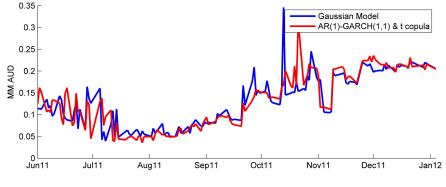
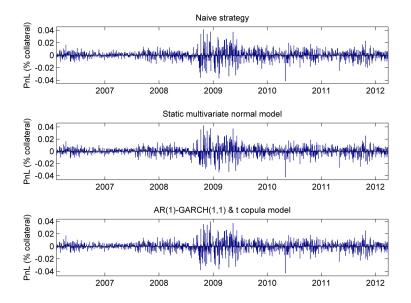
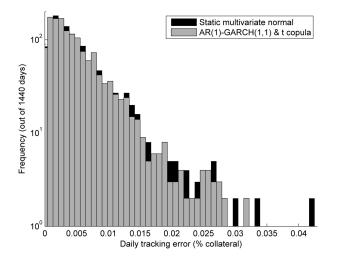


Figure : Optimal posted collateral in JPY, June 2011 - January 2012

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Results

	Naive	MV Gaussian	AR-GARCH & t-
			copula
Avg. daily tracking error (% collateral)	0.54	0.55	0.55
# of margin calls (out of 1440 days)	104	94	71
Frequency of margin call	0.0722	0.0653	0.0493

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Table : Objective: Minimize the daily tracking error while keeping the probability of a margin call under 0.05

Results

	Naive	MV Gaussian	AR-GARCH & t-
			copula
Realized daily VaR (% collateral)	1.19	1.21	1.24
# of margin calls (out of 1440 days)	104	105	78
Frequency of margin call	0.0722	0.0729	0.0542

Table : Objective: Minimize the Value-at-Risk while keeping the probability of a margin call under 0.05

Results

	Naive	MV Gaussian	AR-GARCH & t-
			copula
Avg. daily tail loss (% collateral)	-1.83	-1.85	-1.85
# of margin calls (out of 1440 days)	104	100	71
Frequency of margin call	0.0722	0.0694	0.0493

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Table : Objective: Minimize the Tail Conditional Expectation while keeping the probability of a margin call under 0.05

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Copula-based GARCH:

- are amongst the best model available for multivariate financial time series
- have absolute goodness-of-fit tests now available (parametric bootstrapping)
- provides for better and more robust portfolio engineering and risk management

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